# 9. Recurrent and Transient States

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#### 9.1 Definitions

Define 
$$f_{ii}^{(n)} = P\{X_n = i, X_1 \neq i, \dots, X_{n-1} \neq i | X_0 = i\}$$

= Probability of first recurrence to i is at the  $n^{th}$  step.

$$f_i = f_{ii} = \sum_{n=1}^{\infty} f_{ii}^{(n)}$$
 = Prob. of recurrence to *i*.

<u>Def.</u> A state *i* is <u>recurrent</u> if  $f_i = 1$ .

<u>Def.</u> A state *i* is <u>transient</u> if  $f_i < 1$ .

Define  $T_i$  = Time for first visit to *i* given  $X_0$  = 1. This is the same as Time to first visit to *i* given  $X_k = i$ . (Time homogeneous)

 $\sim$ 

$$m_i = E(T_i | X_0 = i) = \sum_{n=1}^{\infty} n f_{ii}^{(n)}$$
 = mean time for recurrence

<u>Note:</u>  $f_{ii}^{(n)} = P\{T_i = n | X_0 = i\}$ 

Similarly we can define

$$f_{ij}^{(n)} = P\{X_n = j, X_1 \neq j, \dots, X_{n-1} \neq j | X_0 = i\}$$

= Prob. of reaching state j for first time in n steps starting from  $X_0 = i$ .  $f_{ij} = \sum_{n=1}^{\infty} f_{ij}^{(n)} =$  Prob. of ever reaching j starting from i. Consider  $f_{ii} = f_i =$  prob. of ever returning to i. If  $f_i < 1$ ,  $1 - f_i =$  prob. of never returning to i. i.e.

$$1 - f_i = P\{T_i = \infty | X_0 = i\}$$
$$f_i = P\{T_i < \infty | X_0 = i\}$$

<u>TH.</u> If N is no. of visits to  $i|X_0 = i \Rightarrow E(N|X_0 = i) = 1/(1 - f_i)$ <u>Proof:</u>  $E(N|X_0 = i) = E[N|T_i = \infty, X_0 = i]P\{T_i = \infty|X_0 = i\}$   $+E[N|T_i < \infty, X_0 = i]P\{T_i < \infty|X_0 = i\}$  $E(N|X_o = i) = 1 \cdot (1 - f_i) + f_i[1 + E(N|X_0 = i)]$ 

If  $T_i = \infty \Rightarrow$  except for n = 0  $(X_0 = i)$ , there will never be a visit to *i*i.e.  $E(N|T_i = \infty, X_0 = i) = 1$ . If  $T_i < \infty$ , there is sure to be one visit, say at  $X_k$   $(X_k = i)$ . But then  $E(N|T_i < \infty, X_k = i) = E(N|T_i < \infty, X_0 = i)$  by Markov property; i.e.

$$E[N|T_{i} < \infty, X_{0} = i] = 1 + E[N|X_{0} = i]$$
  

$$\therefore E[N|X_{0} = i] = 1 \cdot (1 - f_{i}) + \{1 + E[N|X_{0} = i]\} \cdot (f_{i})$$
  

$$\Rightarrow E[N|X_{0} = i] = 1/(1 - f_{i})$$
  
nother expression for  $E[N|X_{0} = i] = \sum_{n=0}^{\infty} p_{ii}^{(n)}$ 

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#### **Relation to Geometric Distribution**

Suppose *i* is transient  $(f_i < 1)$  and  $N_i = \text{no. of visits to } i$ .

$$P\{N_i = k+1 | X_0 = i\} = f_i^k (1 - f_i), \quad k = 0, 1, \dots$$
$$E(N_i | X_0 = i) = \sum_{k=0}^{\infty} (k+1) f_i^k (1 - f_i) = \sum_{k=0}^{\infty} k f_i^k (1 - f_i) + 1$$

Since

$$(1 - f_i)^{-1} = \sum_{k=0}^{\infty} f_i^k$$
  

$$(1 - f_i)^{-2} = \frac{d}{df_i} (1 - f_i)^{-1} = \sum_{k=0}^{\infty} k f_i^{k-1}$$
  

$$E(N_i | X_o = i) = f_i (1 - f_i) (1 - f_i)^{-2} + 1$$
  

$$= f_i (1 - f_i)^{-1} + 1 = 1/(1 - f_i)$$

TH. 
$$E[N|X_0 = i] = \sum_{n=0}^{\infty} p_{ii}^{(n)}$$
  
Proof. Let  $Y_n = \begin{cases} 1 & \text{if } X_n = i \\ 0 & \text{otherwise} \end{cases}$ 

$$N = \sum_{n=0}^{\infty} Y_n$$

Since  $P\{Y_n = 1 | X_0 = i\} = P\{X_n = i | X_0 = i\} = p_{ii}^{(n)}$ 

$$E(N) = \sum_{n=0}^{\infty} E(Y_n) = \sum_{n=0}^{\infty} p_{ii}^{(n)}$$

E(N) may be finite or infinite

<u>Def:</u> A <u>positive recurrent</u> state is defined by  $f_i = 1, m_i < \infty$ . A <u>null recurrent</u> state is defined by  $f_i = 1, m_i = \infty$ 

Ex: 
$$f_{ii}^{(n)} = \frac{1}{n(n+1)} = \frac{1}{n} - \frac{1}{n+1}$$
  
 $f_i = \sum_{n=1}^{\infty} \left(\frac{1}{n} - \frac{1}{n+1}\right) = 1$ 

But 
$$m_i = \sum_{n=1}^{\infty} n f_{ii}^{(n)} = \sum_{1}^{\infty} \frac{n}{(n+1)n} = \sum_{n=1}^{\infty} \frac{1}{n+1} = \infty$$
 as series does not converge.

# **Classification of States**

	$f_i$	$m_i$
Positive recurrent state	1	$<\infty$
Null recurrent state	1	$\infty$
Transient	< 1	$<\infty$

where  $m_i$  = Expected no. of visits to *i* given  $X_0 = i$ .

In addition the recurring and transient states may be characterized by being periodic or aperiodic.

A state is ergodic if it is aperiodic and positive recurrent.

9.2 Relations Between  $f_i$  and  $p_{ii}^{(n)}$ 

Consider  $p_{ii}^{(n)}$ . Starting from  $X_0 = i$ , the first recurrence to i may be at k = 1, 2, ..., n. Consider the first visit is at time k and at  $X_n, X_n = i$  another visit is made. This probability is  $f_{ii}^{(k)}p_{ii}^{(n-k)}$ . Summing over all k results in

(\*) 
$$p_{ii}^{(n)} = \sum_{k=1}^{n} f_{ii}^{(k)} p_{ii}^{(n-k)}$$

where  $p_{ii}^{(o)} = 1 = P\{X_0 = i | X_0 = i\}$ 

Multiplying (\*) by  $s^n$  and summing

$$\sum_{n=1}^{\infty} p_{ii}^{(n)} s^n = \sum_{n=1}^{\infty} \sum_{k=1}^{n} f_{ii}^{(k)} p_{ii}^{(n-k)} s^n = \sum_{k=1}^{\infty} f_{ii}^{(k)} s^k \sum_{n=k}^{\infty} p_{ii}^{(n-k)} s^{n-k}$$

$$\sum_{n=1}^{\infty} p_{ii}^{(n)} s^n = \sum_{n=1}^{\infty} \sum_{k=1}^{n} f_{ii}^{(k)} p_{ii}^{(n-k)} s^n = \sum_{k=1}^{\infty} f_{ii}^{(k)} s^k \sum_{n=k}^{\infty} p_{ii}^{(n-k)} s^{n-k}$$

$$\boxed{P_{ii}(s) - 1 = F_{ii}(s) P_{ii}(s)}$$
where  $P_{ii}(s) = \sum_{n=0}^{\infty} p_{ii}^{(n)} s^n$ ,  $F_{ii}(s) = \sum_{n=1}^{\infty} f_{ii}^{(n)} s^n$ 

$$\boxed{P_{ii}(s) = \frac{1}{1 - F_{ii}(s)}}$$

Note:

$$\lim_{s \to 1} F_{ii}(s) = F_{ii}(1) = \sum_{n=1}^{\infty} f_{ii}^{(n)} = f_i$$
$$\lim_{s \to 1} F'_{ii}(s) = F'_{ii}(1) = \sum_{n=1}^{\infty} n f_{ii}^{(n)} = m_i$$

#### **Theorems**

(a) If 
$$P_{ii}(1) = \sum_{0}^{\infty} p_{ii}^{(n)} = \infty \Rightarrow f_i = 1.$$
  
Conversely if  $f_i = 1$ ,  $P_{ii}(1) = \infty$   
(b) If  $P_{ii}(1) = \sum_{0}^{\infty} p_{ii}^{(n)} < \infty \Rightarrow f_i < 1.$   
Conversely if  $f_i < 1$ ,  $P_{ii}(1) < \infty$ 

9.3 Limiting Theorems for Generating Functions

Consider 
$$A(s) = \sum_{n=0}^{\infty} a_n s^n$$
,  $|s| \le 1$  with  $a_n \ge 0$ .

1. 
$$\lim_{n \to \infty} \sum_{k=0}^{n} a_k = \lim_{s \to 1} A(s)$$
 where  $s \to 1$  means  $s \to 1-$ .

<u>2.</u> Define

$$a^{*}(n) = \sum_{k=0}^{n} a_{k}/(n+1)$$
$$\lim_{n \to \infty} a^{*}(n) = \lim_{s \to 1} (1-s)A(s)$$

3. Cesaro Limit

The Cesaro limit is defined by  $\lim_{n \to \infty} a^*(n)$  If the sequence  $\{a_n\}$  has a limit  $\Pi = \lim_{n \to \infty} a_n$  then  $\lim_{n \to \infty} a^*(n) = \Pi$ .

The Cesaro limit may exist without the existence of the ordinary limit.

# <u>Ex.</u> $a_n$ : 0, 1, 0, 1, 0, 1, ...

 $\lim_{n \to \infty} a_n \text{ does not exist.}$ 

However

$$a^*(n) = \begin{cases} \frac{1}{2} & \text{if } n \text{ even} \\\\ \frac{1}{2}(1 - \frac{1}{n}) & \text{if } n \text{ is odd} \\\\\\ \lim_{n \to \infty} a^*(n) = \frac{1}{2} \end{cases}$$

#### 9.4 Application to Markov Chains

Consider 
$$p_{ii}^*(n) = \sum_{k=0}^n \frac{p_{ii}^{(k)}}{n+1}$$

 $\sum_{k=0}^{n} p_{ii}^{(k)}$  is expected no. of visits to *i* starting from  $X_0 = i \ (p_{ii}^0 = 1)$ .

Dividing by (n+1),  $p_{ii}^*(n)$  is expected no. of visits per unit time.

<u>Ex.</u> n = 29 days,  $p_{ii}^*(29) = 2/30$ ; i.e. 2 visits per 30 days or 1 visit per 15 days. One would expect mean time between visits = 15 days.

TH. 
$$\lim_{n \to \infty} \sum_{k=0}^{n} \frac{p_{ii}^{(k)}}{n+1} = \frac{1}{m_i}, \text{ where } m_i = \text{expected no. of visits and}$$
$$f_i = 1$$
Proof: Consider  $P_{ii}(s) = \frac{1}{1 - F_{ii}(s)}$ 
$$\lim_{s \to 1} (1-s)P_{ii}(s) = \lim_{n \to \infty} p_{ii}^*(n) = \lim_{s \to 1} \frac{(1-s)}{1 - F_{ii}(s)}.$$
Since  $F_i(1) = f_i$  if  $f_i = 1$  the respected no. of visits and

Since  $F_{ii}(1) = f_i$ , if  $f_i = 1$ , the r.h.s. is indeterminate. Using L'Hopital's rule

$$\lim_{n \to \infty} p_{ii}^*(n) = \frac{1}{F_{ii}'(1)} = \frac{1}{m_i}$$

Recall a <u>positive recurrent</u> state has  $m_i < \infty \implies \lim_{n \to \infty} p_{ii}^*(n) > 0$ 

A <u>null recurrent</u> state has  $m_i = \infty$ 

$$\Rightarrow \lim_{n \to \infty} p_{ii}^*(n) = 0 \text{ or } \lim_{n \to \infty} p_{ii}^{(n)} = 0$$

9.5 Relations Between  $f_{ij}$  and  $p_{ij}^{(n)}$   $(i \neq j)$ 

$$f_{ij}^{(n)} = P\{X_n = j, X_r \neq j, r = 1, 2, \dots, n-1 | X_0 = i\}$$

= Prob. of starting from i and reaching j for first time at  $n^{th}$  step.

$$f_{ij} = \sum_{n=1}^{\infty} f_{ij}^{(n)} \quad i \neq j$$

Proceeding as before  $(i \neq j)$ 

$$p_{ij}^{(n)} = f_{ij}^{(1)} p_{jj}^{(n-1)} + f_{ij}^{(2)} p_{jj}^{(n-1)} + \dots + f_{ij}^{(n)}$$
$$= \sum_{k=1}^{n} f_{ij}^{(k)} p_{jj}^{(n-k)} \quad (p_{jj}^{(0)} = 1)$$

Multiplying by  $s^n$  and summing over n

$$\sum_{n=1}^{\infty} p_{ij}^{(n)} s^n = \sum_{n=1}^{\infty} \sum_{k=1}^{n} f_{ij}^{(k)} p_{jj}^{(n-k)} s^n$$

$$P_{ij}(s) = \sum_{k=1}^{\infty} f_{ij}^{(k)} s^k \sum_{n=k}^{\infty} p_{jj}^{(n-k)} s^{n-k}$$

$$\boxed{P_{ij}(s) = F_{ij}(s) P_{jj}(s)} \quad i \neq j$$

$$\lim_{s \to 1} (1-s) P_{jj}(s) = \lim_{n \to \infty} p_{jj}^*(n) = \lim_{s \to 1} (1-s) P_{ij}(s) / F_{ij}(1)$$

$$\lim_{n \to \infty} p_{jj}^*(n) = \lim_{n \to \infty} \frac{p_{ij}^*(n)}{F_{ij}(1)} = \frac{1}{m_i} \text{ or } \boxed{\lim_{n \to \infty} p_{ij}^*(n) = \frac{F_{ij}(1)}{m_i}}$$

Also 
$$P_{ij}(1) = \sum_{n=1}^{\infty} p_{ij}^{(n)} = F_{ij}(1) \sum_{n=0}^{\infty} p_{jj}^{(n)}.$$
  
Hence if  $\sum_{n=0}^{\infty} p_{jj}^{(n)} = \infty \Rightarrow \sum_{n=0}^{\infty} p_{ij}^{(n)} = \infty \quad (p_{ij}^{(0)} = 0).$   
Similarly if  $\sum_{n=0}^{\infty} p_{jj}^{(n)} < \infty \Rightarrow \sum_{n=0}^{\infty} p_{ij}^{(n)} < \infty$ 

$$\underline{Summary}$$

$$\underline{Transient} \qquad \sum_{n=0}^{\infty} p_{ii}^{(n)} < \infty, \ f_i < 1, \ m_i = 1/(1 - f_i)$$

$$\lim_{n \to \infty} p_{ii}^{(n)} = 0, \ \sum_{n=1}^{\infty} p_{ij}^{(n)} < \infty, \ \lim_{n \to \infty} p_{ij}^{(n)} = 0$$

$$\underline{Positive Recurrent} \qquad \sum_{n=0}^{\infty} p_{ii}^{(n)} = \infty, \ f_i = 1, \ m_i < \infty$$

$$\lim_{n \to \infty} p_{ii}^*(n) > 0 \ (= 1/m_i)$$

$$\lim_{n \to \infty} p_{ij}^*(n) > 0 \ (= F_{ij}(1)/m_i)$$

$$\underline{Negative Recurrent} \qquad \sum_{n=0}^{\infty} p_{ii}^{(n)} = \infty, \ f_i = 1, \ m_i = \infty$$

$$\lim_{n \to \infty} p_{ii}^*(n) = 0, \ \lim_{n \to \infty} p_{ii}^{(n)} = 0$$

# 9.6 Periodic Processes

Suppose transition probabilities have period d. Then

$$p_{ij}^{(n)} = 0, \ p_{ii}^{(n)} = 0 \ \text{if } n \neq rd \ r = 1, 2, \dots$$
$$p_{ij}^{(n)} \ge 0, \ p_{ii}^{(n)} \ge 0 \ \text{if } n = rd$$
$$P_{ii}(s) = \sum_{r=0}^{\infty} p_{ii}^{(rd)} s^{rd} = \sum_{r=0}^{\infty} p_{ii}^{(rd)} z^{r}, \ z = s^{d}$$
$$F_{ii}(s) = \sum_{r=1}^{\infty} f_{ii}^{(rd)} s^{rd} = \sum_{r=1}^{\infty} f_{ii}^{(rd)} z^{r}$$

We now have a power series in z

$$\lim_{z \to 1} P_{ii}(Z) = \sum_{r=0}^{\infty} p_{ii}^{(rd)}$$
$$\lim_{z \to 1} (1-z) P_{ii}(z) = \sum_{n \to \infty} \sum_{r=0}^{n} \frac{p_{ii}^{(rd)}}{n+1}$$

Note: 
$$E(N|X_0 = i) = \sum_{1}^{\infty} n f_{ii}^{(n)} = d \sum_{r=1}^{\infty} r f_{ii}^{(rd)} = m_i$$

However 
$$F'_{ii}(1) = \sum_{r=1}^{\infty} f_{ii}^{(rd)} r = m_i/d$$

Since  $P_{ii}(Z) = 1/[1 - F_{ii}(z)]$ 

$$\lim_{n \to \infty} p_{ii}^*(n) = d/m_i \text{ or } \lim_{n \to \infty} p_{ii}^{(nd)} = d/m_i$$

#### 9.7 Closed Sets

<u>Def.</u> A set of states C is closed if no state outside C can be reached from any state in C; i.e.,  $p_{ij} = 0$  if  $i \in C$  and  $j \notin C$ .

Absorbing state: Closed set consisting of a single state.

<u>Irreducible Chain:</u> If only closed set is the set of all states. (Every state can be reached from any other state).

This means that we can study the behavior of states in C by omitting all other states.

<u>Th.</u>  $i \leftrightarrow j$ , *i* is recurrent  $\Rightarrow j$  recurrent  $i \leftrightarrow j, i \text{ is transient} \Rightarrow j \text{ transient}$  $\sum_{r=0}^{\infty} p_{jj}^{(r)} \ge \sum_{r=0}^{\infty} p_{jj}^{(r+n+m)} = \sum_{r=0}^{\infty} \sum_{k \in S} p_{jk}^{(m)} p_{kk}^{(r)} p_{kj}^{(n)}$  $\geq \sum_{i=1}^{\infty} p_{ji}^{(m)} p_{ii}^{(n)} p_{ij}^{(n)} = p_{ji}^{(m)} p_{ij}^{(n)} \sum_{i=1}^{\infty} p_{ii}^{(r)}$ Thus if  $\sum_{i=1}^{\infty} p_{ii}^{(r)} = \infty$ ,  $\sum_{i=1}^{\infty} p_{jj}^{(r)} = \infty$ Suppose i is transient,  $i \leftrightarrow j$  and assume j is recurrent. By theorem just proved  $\underline{i}$  then must be recurrent. However this is a contradiction

 $\Rightarrow \underline{j}$  is transient.

Summary (Aperiodic, irreducible)

1. If  $i \leftrightarrow j$  and j is positive recurrent  $\Rightarrow i$  is positive recurrent

$$\lim_{n \to \infty} p_{ij}^{(n)} = \lim_{n \to \infty} p_{jj}^{(n)} = \prod_j = 1/m_j.$$

2. If  $i \leftrightarrow j$  and j is null recurrent  $\Rightarrow i$  is null recurrent

$$\lim_{n \to \infty} p_{jj}^{(n)} = 0, \quad \lim_{n \to \infty} p_{ij}^{(n)} = 0$$

or 
$$P^{(\infty)} = \lim_{n \to \infty} P^{(n)} = \lim_{n \to \infty} P^n = 0.$$

3. If  $i \leftrightarrow j$  and j is transient  $\Rightarrow i$  is transient

$$\lim_{n \to \infty} p_{jj}^{(n)} = 0, \lim_{n \to \infty} p_{ij}^{(n)} = 0$$

# 9.8 Decomposition Theorem

(a) The states of a Markov Chain may be divided into two sets (one of which may be empty). One set is composed of all the recurring states, the other of all the transient states.

(b) The recurrent states may be decomposed uniquely into two closed sets. Within each closed set all states inter-communicate and they are all of the same type and period. Between any two closed sets no communication is possible.

Ex. Decc	ompositio	n of a	Fini	ite C	Chain										
P =		$C_0$			$C_1$				$C_2$				$C_3$		
	$C_0$	1 0	• •	0	•••	0	• •	0	•••	0	• •	0	•••	0	
		0 0	• •	0	•••	0	• •	0	•••	0	• •	0	•••	0	
	$C_1$	0	• •		$P_1$		• •		0		• •		0		
	$C_2$	0	• •		0		• •		$P_1$		• •		0		
	$C_3$	A	• •		B		• • •		C		• •		D		

- $C_0$ : Consists of two absorbing states
- $C_1$ : Consists of closed recurrent states
- $C_2$ : Consists of closed recurrent states
- $C_3$ : Consists of transient states

A: Transitions  $C_3 \to C_0$  B: Transitions  $C_3 \to C_1$ 

C: Transitions  $C_3 \rightarrow C_2$  D: Transitions  $C_3 \rightarrow C_3$ 

- 9.9 <u>Remarks on Finite Chains</u>
  - 1. A finite chain cannot consist only of transient states.

If i, j are transient  $\lim_{n \to \infty} p_{ij}^{(n)} = 0$ , However

$$\sum_{j \in S} p_{ij}^{(n)} = 1$$

leading to a contradiction as  $n \to \infty$ .

2. A finite chain cannot have any null recurrent states.

The one step transition probabilities within a closed set of null recurrent states form a stochastic matrix P such that  $P^n \to 0$  as  $n \to \infty$ . This is impossible as  $\sum_{i \in S} p_{ij}^{(n)} = 1$ .

#### 9.10 Perron-Frobenius Theorem

Earlier we had seen that if P has a characteristic root (eigenvalue) = 1 of multiplicity 1, and all other  $|\lambda_i| < 1$ , then

$$P^{(\infty)} = E_1.$$

The conditions under which this is true are proved by the Perron-Frobenius Theorem. The necessary and sufficient conditions are:

- P: aperiodic
- P: positive recurrent  $(m_i < \infty)$ .

Then  $P_1^{(\infty)} = P^{\infty} = E_1 = 1y'$ where y'P = y' and  $\underline{1} = \begin{pmatrix} 1 & 1 & \cdots & 1 \end{pmatrix}'$ <u>Def.</u> A state is ergodic if it is aperiodic and positive recurrent. 9.11 Determining Recurrence and Transience when Number of States is Infinite

Compute  $f_i = P\{T_i < \infty | X_0 = i\}$  in a closed communicating class.

 $\Rightarrow$  All states are recurrent if  $f_i = 1$ ; All states are transient if  $f_i < 1$ .

Ex. 
$$S = \{0, 1, 2, ...\}, \quad p_{i,0} = q_i \quad p_{i,i+1} = p_i,$$
  

$$X_{n+1} = \begin{cases} 0 & \text{with } q_i \\ X_n + 1 & \text{with } p_i \end{cases}$$

$$P\{T_0 > n | X_0 = 0\} = P\{X_1 = 1, X_2 = 1, ..., X_n = n | X_0 = 0\}$$

$$= \prod_{i=0}^{n-1} p_i$$

$$P\{T_0 < \infty | X_0 = 0\} = 1 - \lim_{n \to \infty} P\{T_0 > n | x_0 = 0\} = 1 - \prod_{i=0}^{\infty} p_i$$

$$\therefore \text{ State } 0 \text{ (and all states in closed class) are recurrent iff } \prod_{i=0}^{\infty} p_i = 0$$

 $\therefore$  State 0 (and all states in closed class) are recurrent iff  $\prod_{i=0}^{\infty} p_i = 0$ 

Ex. Random Walk on Integers

$$S = \{0, \pm 1, \pm 2, \dots\}$$

$$p_{i,i+1} = p, \quad p_{i,i-1} = q, \quad p+q = 1$$

 $p_{00}^{(2n+1)} = 0, \ n \ge 0$  (Go from 0 to 0 in odd number of transitions)

$$p_{00}^{(2n)} = \binom{2n}{n} p^n q^n$$

Consider

$$\sum_{n=1}^{\infty} p_{00}^{(n)} = \sum_{n=1}^{\infty} p_{00}^{(2n)} = \sum_{n=1}^{\infty} \frac{(2n)!}{n!n!} p^n q^n$$

Note: Ratio Test: If 
$$A = \sum_{n=0}^{\infty} a_n$$
 and  
 $\frac{a_{n+1}}{a_n} < 1$  series converges as  $n \to \infty$   
 $\frac{a_{n+1}}{a_n} > 1$  series diverges as  $n \to \infty$   
 $\frac{p_{00}^{2n+2}}{p_{00}^{2n}} = \frac{(2n+1)(2n+2)}{(n+1)(n+1)}pq \to 4pq$  as  $n \to \infty$ 

If 
$$p \neq q \quad 4pq < 1$$
.  
If  $p = q = 1/2 \quad 4pq = 1$  and test is inconclusive.  
 $\binom{2n}{s}p^sq^{2n-s} \sim N(2np, 2npq) = N(n, n/2)$  if  $p = 1/2$   
 $\sim e^{-(s-n)^2/n} / \sqrt{2\pi \frac{n}{2}}$ 

Since 
$$p_{00}^{(2n)} \sim \frac{1}{\sqrt{\pi n}}$$
  
$$\sum_{n=1}^{\infty} p_{00}^{(2n)} \cong \sum_{n=1}^{\infty} \frac{1}{\sqrt{\pi n}} \text{ series diverges}$$

. State 0 is recurrent.

<u>Th.</u> An irreducible Markov Chain with  $S = \{0, 1, 2, ...\}$  and Transition Prob.  $\{p_{ij}\}$  is transient iff

$$y_i = \sum_{j=1}^{\infty} p_{ij} y_j$$
  $i = 1, 2, \dots$ 

has a non-zero bounded solution

<u>Proof: P.88</u>

<u>Ex.</u> Random walk on  $S = \{0, 1, 2, ...\}$ 

$$p_{i,i+1} = p_i, \ p_{i,i-1} = q_i \ p_{i,i} = r_i, \ q_0 = 0 \ (p_i + q_i + r_i = 1)$$

Equations:

$$y_1 = p_{11}y_1 + p_{12}y_2 = r_1y_1 + p_1y_2 \Rightarrow y_2 = \left(1 + \frac{q_1}{p_1}\right)y_1$$

 $y_2 = p_{21}y_1 + p_{22}y_2 + p_{23}y_3 = q_2y_1 + r_2y_2 + p_2y_3$ 

$$\Rightarrow y_3 = \left(1 + \frac{q_1}{p_1} + \frac{q_1q_2}{p_1p_2}\right)y_1$$

In general,

$$y_n = \left[1 + \sum_{k=1}^{n-1} \frac{q_1 q_2 \dots q_k}{p_1 p_2 \dots p_k}\right] y_1, \ n \ge 1$$

$$y_n = \left(\sum_{k=0}^{n-1} \alpha_k\right) y_1, \quad \alpha_k = \frac{q_1 q_2 \dots q_k}{p_1 p_2 \dots p_k}, \ \alpha_0 = 1$$

Thus the solution is bounded if  $\sum_{k=0}^{\infty} \alpha_k < \infty$ 

#### 9.12 Revisiting Statistical Equilibrium

Assume the Markov Chain has all states which are irreducible positive recurrent aperiodic. (Called Ergodic Chain).

Earlier we had shown

$$\lim_{n \to \infty} p_{ij}^{(n)} = \lim_{n \to \infty} p_{jj}^{(n)} = 1/m_j = \Pi_j$$

 $\Rightarrow \Pi_j$  is given by the solution

$$\Pi_j = \sum_{i \in S} \Pi_i p_{ij} \quad \text{where} \quad \sum_{j \in S} \Pi_j = 1$$

or in matrix notation we can write the linear equations as  $\Pi = P\Pi$  where  $\Pi : k \times 1$ ,  $P : k \times k$ .

<u>Proof</u>  $a_j(n) = P\{X_n = j\}$  and we will show that  $\lim_{n \to \infty} a_j(n) = \prod_j$  $a_j(n+m) = \sum_{i \in S} \prod_i p_{ij}^{(n)}$ 

Take 
$$m \to \infty$$
  $\Pi_j = \sum_{i \in S} \Pi_i p_{ij}^{(n)}$  for any  $n$   
If  $n = 1$   $\Pi_j = \sum_{i \in S} \Pi_i p_{ij}$   
Allow  $n \to \infty$   $\Pi_j = \left(\sum_{i \in S} \Pi_i\right) \Pi_j$   
which is true if  $\sum_{i \in S} \Pi_i = 1$   
In the above we made use of

$$\lim_{n \to \infty} p_{ij}^{(n)} = \lim_{n \to \infty} p_{jj}^{(n)} = \Pi_j$$

which holds for ergodic chains.

If chain was transient or null recurrent

$$p_{ij}^{(n)} = p_{jj}^{(n)} \to 0 \text{ as } n \to \infty \text{ and } \Pi_j = \sum_{i \in S} \Pi_i p_{ij}^{(n)} \to 0 \text{ as } n \to \infty,$$

and result does not hold.

<u>Th:</u> If  $a_0(j) = P\{X_0 = j\} = \prod_j$  then  $P\{a_n = j\} = \prod_j$  for all n.

This theorem states that if the initial probabilities correspond to the limiting probabilities, then for any  $n \quad p\{X_n = j\} = \prod_j$ .

#### Proof:

In general  $a_j(n) = \sum_{i \in S} a_0(i) p_{ij}^{(n)}$  and in matrix notation we can write  $a_n = P^n a_0$ 

If  $a_o = \Pi$ ,  $a_n = P^n \Pi$ . But  $\Pi$  is defined by  $\Pi = P \Pi$ .  $\Rightarrow a_n = P \Pi = \Pi$ 

This is the reason why  $\Pi$  is sometimes referred to as the *stationary distribution*.

#### 9.13 Appendix. Limit Theorems for Generating Functions

<u>Definition</u>: Let  $A(z) = \sum_{n=0}^{\infty} a_z^n$  denote a power series. In our application z will always be real, however all the results also hold if z is a complex number.

<u>Theorem</u> If  $\{a_n\}$  are bounded, say  $|a_n| \leq B$ , then A(z) converges for at least |z| < 1.

Proof: 
$$A(z) = |A(z)| \le \sum_{n=0}^{\infty} |a_n| z^n | \le B \sum_{n=0}^{\infty} |z|^n = B/1 - |z|$$

<u>Definition</u>: The number R is called the <u>radius of convergence</u> of the power series A(z) if A(z) converges for |z| > R. Without loss of generality the radius of convergence can be taken as R = 1. Note if R is the radius of convergence of A(z) we can write  $A(z) = \sum_{0}^{\infty} b_n y^n$  and the radius of convergence of y will be unity.

# Properties of A(z)

(i) If R is the radius of convergence

$$R^{-1} = \lim_{n \to \infty} \sup(a_n)^{1/n}$$

(ii) Within the interval of convergence (-R < z < R), A(z) has derivatives of all orders which may be obtained by term-wise differentiation. Similarly the integral  $\int_{a}^{b} A(z)dz$  is given by term-wise integration for any (a, b) in (-R, R). (iii) If A(z) and  $B(z) = \sum_{0}^{\infty} b_n z^n$  both converge and are equal for all |z| < R, then  $a_n = b_n$ . (iv) No general statement can be made about the convergence of the series on the boundary |z| = R; i.e.  $\sum_{n} a_n R^n$  may or may not be finite. Two important theorems on power series are:

<u>Abel's Theorem</u> Suppose A(z) has a radius of convergence R = 1 and  $\sum_{0}^{\infty} a_n$  is convergent to s. Then

$$\lim_{z \to 1-} A(z) = \sum_{n=0}^{\infty} a_n$$

If the coefficients  $\{a_n\}$  are non-negative, the result continues to hold whether or not the sum on the right is convergent. Note:

$$A_N(z) = \sum_{n=0}^N a_n z^n = \sum_{n=0}^N (s_n - s_{n-1}) z^n, \qquad s_n = \sum_{i=0}^n a_i$$
$$= \sum_{n=0}^N s_n z^n - z \sum_{n=0}^{N-1} s_n z^n$$
$$= (1 - z) \sum_{n=0}^N s_n z^n + s_N z^N$$

 $\lim_{z \to 1^-} A_n(z) = s_N$  and taking the limit as  $N \to \infty$ 

$$\lim_{N \to \infty} s_N = s$$

<u>Theorem</u>: If the sequence  $\{b_n\}$  converges to a limit  $b (\lim_{n \to \infty} b)$ , then

$$\lim_{z \to 1-} (1-z) \sum_{n=0}^{\infty} b_n z^n = b$$

Proof: 
$$(1-z)\sum_{\substack{n=0\\N}}^{\infty} b_n z^n = \sum_{\substack{n=0\\n=0}}^{\infty} (b_n - b_{n-1}) z^n \quad (b_{-1} = 0)$$

We can write 
$$\sum_{n=0}^{\infty} (b_n - b_{n-1}) z^n = (1-z) \sum_{n=0}^{\infty} s_n z^n + s) n z^N$$

where

$$s_n = \sum_{i=0}^n (b_i - b_{i-1}) = b_0 + (b_1 - b_0) + (b_2 - b_1) + \ldots + (b_n - b_{n-1}) = b_n$$

Therefore 
$$(1-z)\sum_{n=0}^{N} b_n a^n = (1-z)\sum_{n=0}^{N} s_n z^n + b_n z^N$$
  
and  $\lim_{z \to 1^-} (1-z)\sum_{n=0}^{N} b_n z^n = b_n$ , so that as  $N \to \infty$ ,  $\lim_{N \to \infty} b_N$ 

n=0

= b.